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Learning on the Job and the Cost of Business Cycles*

Karl Walentin[†] and Andreas Westermark[‡]

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Abstract

We show that business cycles reduce welfare through a decrease in the average level of employment and output in a labor market search model with learning on-the-job and skill loss during unemployment. A negative correlation between unemployment and vacancies (i.e. a Beveridge curve) combined with a matching function implies that business cycles tend to reduce the average number of new jobs and employment. Then, since learning on-the-job implies that aggregate human capital is increasing in employment, it follows that aggregate volatility reduces human capital. This, in turn, reduces the incentives to post vacancies, further reducing employment and human capital. We quantify this mechanism using a carefully calibrated model and find the output and welfare cost of business cycles to be large.

Keywords: Search and matching, labor market, human capital, stabilization policy, skill loss.

JEL classification: E32, J64.

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1 Introduction

A major question in macroeconomics is how large the welfare costs of business cycles are. Since Lucas (1987), it has been well established that the cost of aggregate consumption fluctuations is negligible. Business cycles can induce welfare costs in other ways though, e.g. through their effect on the cross-sectional distribution of consumption (Imrohoroğlu, 1989, and many others). Furthermore, business cycles may affect welfare negatively by reducing the average level of output, a view that has been argued by DeLong and Summers (1989), Hassan and Mertens (2017) and Summers (2015). Another strand of the literature highlights the effect of human capital dynamics on macroeconomic fluctuations, see e.g., Kehoe, Midrigan and Pastorino (2015) and Krebs and Scheffel (2017).

Our paper adds to this literature by presenting a new mechanism for how business cycles reduce the level of output. We show that business cycles substantially reduce the level of employment, output and welfare in a labor market search model with human capital dynamics. The key mechanism of the paper is as follows: It is well established that the Beveridge correlation is negative, i.e. that vacancies and unemployment are negatively correlated in the data (see e.g., Fujita and Ramey, 2012). Via the matching function, this implies that business cycles tend to reduce the average number of new jobs and hence employment. At an intuitive level, this happens because vacancies and therefore job finding rates in general are high when unemployment is low, thereby yielding fewer new jobs than in absence of business cycles.^{1,2} Then, since learning on-the-job and skill loss during unemployment implies that average human capital is increasing in employment, it follows that aggregate volatility reduces human capital. This, in turn, reduces the incentives to post vacancies, further reducing employment and so on in a vicious circle. This amplification mechanism for how aggregate volatility reduces employment,

¹In a simple search and matching model with a standard Cobb-Douglas matching function, the number of new jobs is given by

$$m_t = f_t u_t = \left(\frac{v_t}{u_t} \right)^{1-\omega} u_t.$$

where f denotes the job finding rate and $\omega < 1$ is the matching function elasticity. Clearly, the number of new jobs is a nonlinear function of vacancies (v) and unemployment (u), indicating that volatility matters for the average number of new jobs. Let bars denote variables in absence of aggregate volatility and “ E ” denote the unconditional expectation in an economy with aggregate volatility. Using the employment flow equation $1 - u_t = (1 - \delta)(1 - u_{t-1}) + m_t$ and letting δ denote the exogenous separation rate, we can derive an expression for the change in new jobs induced by aggregate volatility:

$$Em - \bar{m} \approx \frac{\delta}{\delta + \bar{f}} \left\{ (1 - \omega) \left(\frac{\bar{f}}{\bar{v}} \text{cov}(v, u) - \frac{\bar{f}}{\bar{u}} \text{var}(u) \right) + (Ef - \bar{f}) Eu \right\}$$

where we have used the first-order approximation of $\text{cov}(f, u) = (1 - \omega) (\bar{f}/\bar{v} \cdot \text{cov}(v, u) - \bar{f}/\bar{u} \cdot \text{var}(u))$. As can be seen from the expression for $Em - \bar{m}$, the number of new jobs and hence employment is lower under aggregate volatility if the Beveridge correlation is negative (i.e. $\text{cov}(v, u) < 0$) and $Ef \leq \bar{f}$. This result is related to Jung and Kuester (2011) that states conditions on $\text{cov}(f, u)$ and $Ef - \bar{f}$ for when aggregate volatility implies a reduction of employment.

²More generally, any convex cost (or concave benefit or production function) in any cyclical variable tends to induce a negative relationship between aggregate volatility and average consumption or employment. Prominent examples are convex capital adjustment costs and convex vacancy posting costs, both of which are commonly assumed in the business cycle literature.

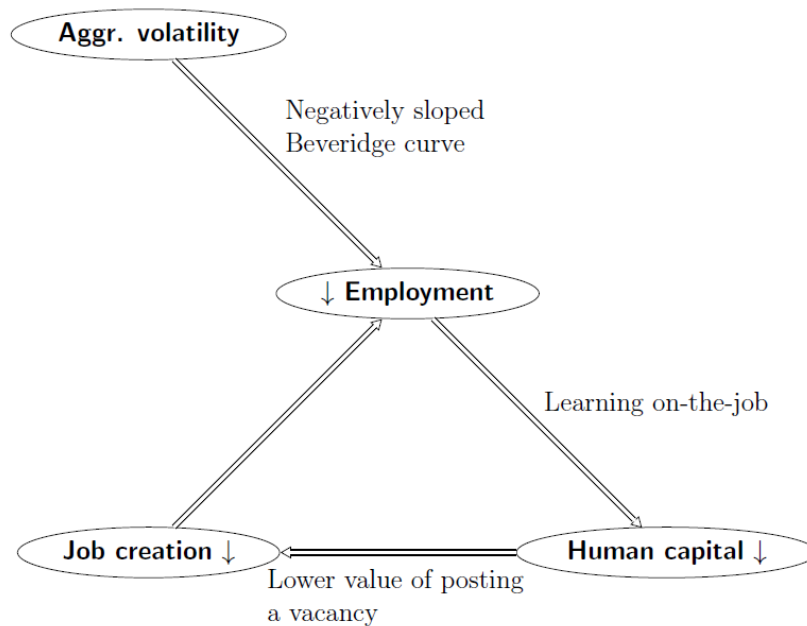


Figure 1: Illustration of main mechanism - how aggregate volatility reduces employment, human capital and thereby output.

human capital and thereby output is illustrated graphically in Figure 1. The size of the cost of business cycles generated by this mechanism is accordingly largely determined by how sensitive the human capital distribution is to changes in employment and how sensitive job creation is to changes in the human capital distribution. Since our mechanism works through the average level of consumption, it is fundamentally different from most of the cost of business cycles literature, which analyses the effects of business cycles on welfare through (aggregate or idiosyncratic) consumption volatility. Our amplification mechanism also extends beyond the cost of business cycles. For example, the effect of a change in taxation or unemployment benefits that affects average employment will be amplified by the human capital mechanism that we have outlined.

We use a search and matching framework with general human capital dynamics (learning on-the-job and skill loss during unemployment) to model the relationship between business cycles and the average level of output. As argued above, an important determinant of the size of the cost of business cycles is how sensitive job creation is to changes in the human capital distribution of both unemployed and employed workers. Thus, we allow for on-the-job search to capture the effect of employed workers' human capital on job creation. In addition, to allow for a flexible bargaining framework in a context with on-the-job search, we use the bargaining protocol from Cahuc, Postel-Vinay and Robin (2006), henceforth CPVR. This framework implies that workers get the value of their outside option plus a

share β of the value of the match above the outside option. We are not aware of any model that uses the bargaining framework of CPVR in a setting with aggregate uncertainty. We propose and implement an algorithm for solving models using global solution methods with on-the-job search and heterogeneous workers and firms under aggregate uncertainty in a context with positive bargaining strength of workers. Thus, the paper also makes a methodological contribution. In our mind, our solution algorithm is useful for future research where heterogeneity in the labor market interacts with the business cycle.

The main goal for our exercise is to provide a credible quantification of the cost of business cycles through the mechanism we have sketched above. One key determinant of this cost is the speed of human capital gain when employed relative to the loss during unemployment. We estimate the human capital gains when employed by matching the empirical “return to experience” (wage profile of employed workers) reported by Buchinsky et al. (2010). The model is calibrated by matching the return to experience and other relevant moments, including volatility of GDP and unemployment, standard worker flow moments and the degree of wage dispersion. We then compute the cost of business cycles by comparing the results for our full model to the results from the same model, but without aggregate volatility. We find that business cycles reduce steady state employment, GDP and welfare by substantial amounts. In particular, eliminating aggregate volatility increases welfare (GDP) by 0.52-1.49 percent (1.45 percent), depending on the interpretation of the flow value of unemployment. These are fairly large effects. Accounting for the transition dynamics, the welfare gains of eliminating business cycles are smaller, 0.20-1.09 percent. Human capital dynamics are pivotal for the results - if we disable them in our model, the implied employment, GDP and welfare losses from business cycles are negligible. Note that, since we assume risk neutral agents and hence abstract from, e.g., the direct welfare costs of consumption volatility, we do not capture the full welfare cost of business cycles and our results can accordingly be interpreted as a lower bound for these costs.

There is indicative empirical support for the relationship between aggregate volatility, unemployment and output implied by our model. Hairault et al. (2010) uses data for 20 OECD countries for the period 1982-2003 and finds significant positive effects of TFP volatility on average unemployment. There is also ample evidence of a significant negative relationship between volatility of output and the average growth rate of output, see e.g., Ramey and Ramey (1995) and Luo et al. (2016). Direct evidence of human capital dynamics, in the form of effects on measurable skills, is documented by Edin and Gustavsson (2008). They find sizeable skill loss effects of unemployment. Additional evidence is provided by Schmieder, von Wachter and Bender (2016). They estimate a substantial casual effect on the re-employment wage of an additional month of unemployment, also indicating considerable loss of

human capital. There is also evidence that local labor market conditions affect future “employability” of workers. Yagan (2017) establishes a strong link between local shocks to employment growth during the Great Recession, 2007-2009, and the 2015 employment rates of workers exposed to these shocks and argues that this is due to depreciation of general human capital during non-employment spells.

Three papers have previously analyzed the effect of business cycles on welfare through the average level of output in a search and matching labor market setting. Den Haan and Sedlacek (2014) quantified the cost of business cycles in a setting where an agency problem generates inefficient job separations in downturns, thereby reducing employment and GDP. Our framework does not include any such agency problem and is bilaterally efficient. Jung and Kuester (2011) quantified the effects on employment and welfare of the negative correlation between the job finding rate and the unemployment rate. They did so in a simpler setting than ours, using a solution method of local second-order approximations, with wages assumed to be independent of labor market tightness.³ Hairault et al. (2010) also studied this issue in a setting without human capital dynamics. Both Jung and Kuester (2011) and Hairault et al. (2010) found substantially smaller effects on GDP and welfare of business cycles than our results indicate. Furthermore, our model also shares mechanisms with a number of papers that analyze earnings losses from displacement (Burdett, Carrillo-Tudela and Coles, 2015, Huckfeldt, 2016, Jarosch, 2015, Jung and Kuhn, 2016, and Krolikowski, 2017).

The paper is outlined as follows. Section 2 presents the model, Section 3 documents the calibration and Section 4 provides the quantitative results. Finally, Section 5 concludes.

2 Model

We set up a business cycle model with a search and matching labor market and human capital dynamics. We allow for on-the-job search to capture the direct effect of employed workers’ human capital on vacancy postings. The basic building blocks of our model are similar to Lise and Robin (2017), henceforth LR, except for the wage bargaining where we follow CPVR.⁴ This wage setting framework implies that workers get the value of their outside option plus a share β , reflecting their bargaining strength, of the value of the match above the outside option. When a worker is hired out of unemployment the outside option is the value of unemployment. If instead an employed worker receives a poaching offer from another firm, the outside option is the value of the second-best match.

³In an extension they allowed for learning on-the-job, but assumed a weaker dependence of human capital on employment than we do. Furthermore, Jung and Kuester did not describe our main mechanism, the vicious circle laid out in Figure 1.

⁴Compared to LR, the features we add are i) positive bargaining power of workers, and ii) learning on the job as well as skill loss during unemployment. A simplification compared to LR is that in our model the match-specific productivity y of a match is not known when a vacancy is posted.

In terms of human capital dynamics, the model is in the tradition of Pissarides (1992) and Ljungqvist and Sargent (1998). As in these papers, we model general human capital as stemming from learning on-the-job and skill loss during unemployment. Worker human capital, denoted by x , follows a stochastic process and $\pi_{xe}(x, x')$ ($\pi_{xu}(x, x')$) denote the Markov transition probability for the worker's human capital level while employed (unemployed).⁵ Firm match-specific productivity is denoted by y .

To summarize the above aspects of our model, in any time period there is heterogeneity across employed workers in terms of human capital x , match-specific productivity y and wage w . Unemployed workers only differ in terms of their human capital.

Utility is linear in consumption and there is no physical capital. Each firm employs (at most) one worker, and output from a match is $p(x, y, z) = xyz$ where z is an aggregate TFP shock with Markov transition probability $\pi(z, z')$.

2.1 Timing

Let us start the detailed model description by providing an overview of the timing protocol. The sequence of events within a period are as follows. First, the aggregate productivity shock z and the idiosyncratic human capital shocks x are realized. Second, a fraction ν of workers die and are replaced by newborn unemployed workers with human capital at the lowest possible level, \underline{x} . Third, separations into unemployment occur. Then, firms post vacancies and workers search for jobs. Finally, new matches are formed, wages are set and production takes place.

2.2 Separations

The ability of recently separated workers to search for jobs within the period, makes it convenient to define match values and match surplus both before and after the search phase has occurred, i.e., at the separation stage and the matching stage. The surplus of a match at the separation stage is $S^s(x, y, z, \Gamma)$ where Γ denotes the endogenous aggregate state. Matches with $S^s(x, y, z, \Gamma) < 0$ are endogenously dissolved. In addition, a fraction δ of matches are exogenously destroyed every period.

The stock of unemployed workers after separations when the aggregate productivity evolves from

⁵Our human capital dynamics differ slightly from Ljungqvist and Sargent (1998, 2008) and Jung and Kuester's (2011) extension with human capital in that we do not assume a sudden loss of general human capital when a worker separates from a job. These papers abstract from heterogeneity in match-specific productivity and presumably therefore assume, as a short-cut, that part of the human capital loss occurs when a worker is separated from a job. This reduces the dependence of the human capital distribution on employment (or any endogenous variable in the model), especially if one only allows for exogenous separations.

z_{-1} to z is:

$$u^s(x, z) = \nu \mathbf{1}\{x = \underline{x}\} + (1 - \nu) \left[\sum_{x_{-1} \in X} u(x_{-1}, z_{-1}) \pi_{xu}(x_{-1}, x) \right. \\ \left. + \sum_{y \in Y} \sum_{x_{-1} \in X} (\mathbf{1}\{S^s(x, y, z, \Gamma) < 0\} + \delta \mathbf{1}\{S^s(x, y, z, \Gamma) \geq 0\}) h(x_{-1}, y, z_{-1}) \pi_{xe}(x_{-1}, x) \right] \quad (1)$$

where $\mathbf{1}\{\cdot\}$ is the indicator function, u (h) is the distribution of unemployed (employed) workers at the end of a period, X is the set of human capital states and Y is the set of match-specific productivities. Here, the first term is the newborn workers and the remaining terms captures the evolution of the surviving workers.

The stock of matches of type (x, y) at this point is:

$$h^s(x, y, z) = (1 - \delta)(1 - \nu) \sum_{x_{-1} \in X} \mathbf{1}\{S^s(x, y, z, \Gamma) \geq 0\} h(x_{-1}, y, z_{-1}) \pi_{xe}(x_{-1}, x). \quad (2)$$

2.3 Search and matching

An employed worker exerts search effort s_1 . The search effort of unemployed workers is normalized to unity. Accordingly, the aggregate amount of search effort is:

$$L \equiv \sum_{x \in X} u^s(x, z) + s_1 \sum_{x \in X} \sum_{y \in Y} h^s(x, y, z). \quad (3)$$

Vacancy posting costs are linear and each vacancy posted incurs a cost of c_0 . The free entry condition for vacancy creation therefore implies:

$$c_0 = qJ(z, \Gamma). \quad (4)$$

where q is the probability of a firm meeting a worker and J is the expected value of a new match for a firm, as defined below.

We assume the following Cobb-Douglas meeting function:

$$M \equiv \min \{ \alpha L^\omega V^{1-\omega}, L, V \} \quad (5)$$

where V is the number of vacancies posted. The probability of a firm meeting a worker (assuming an interior solution) is:

$$q = \frac{M}{V} = \alpha \theta^{-\omega},$$

where $\theta \equiv \frac{V}{L}$ is labor market tightness. Together with the matching function (5), this implies that equilibrium vacancy postings are determined by:

$$V = L \left(\frac{\alpha J(z, \Gamma)}{c_0} \right)^{\frac{1}{\omega}}. \quad (6)$$

We can then write labor market tightness as a function of z and Γ :

$$\theta(z, \Gamma) = \left(\frac{\alpha J(z, \Gamma)}{c_0} \right)^{\frac{1}{\omega}}. \quad (7)$$

Finally, the probability that an unemployed worker meets a firm (the job meeting rate) is, assuming an interior solution:

$$f(z, \Gamma) = \frac{M}{L} = \alpha \theta(z, \Gamma)^{1-\omega}. \quad (8)$$

2.4 Values

A worker who is unemployed during the production phase receives a flow payoff of $b(x, z)$ representing unemployment insurance, utility of leisure and value of home production. The value of unemployment at the matching stage is:

$$\begin{aligned} B(x, z, \Gamma) = & b(x, z) \\ & + \frac{1-\nu}{1+r} \sum_{x' \in X} \sum_{z' \in Z} \left[\sum_{y' \in Y} f(z', \Gamma') [B(x', z', \Gamma') + \beta \max\{P(x', y', z', \Gamma') - B(x', z', \Gamma'), 0\}] g(y') \right. \\ & \left. + (1 - f(z', \Gamma')) B(x', z', \Gamma') \right] \times \pi_{xu}(x, x') \pi(z, z'), \end{aligned} \quad (9)$$

where r is the discount rate, Z is the set of aggregate productivity states, P the value of a match and $g(y)$ is the probability density function (pdf) of the productivity of newly created matches. Thus, B is the flow payoff b plus the job meeting rate $f(z', \Gamma')$ times the discounted value of a job tomorrow plus $(1 - f(z', \Gamma'))$ times the discounted value of being unemployed tomorrow. The max operator ensures that only matches with positive surplus are formed. Note that while a worker is unemployed his human capital (weakly) decreases from x to x' with probability $\pi_{xu}(x, x')$.

The match value at the matching stage, using that the job meeting rate for employed workers is

$s_1 f(z', \Gamma')$, can be written as follows:

$$\begin{aligned}
P(x, y, z, \Gamma) &= p(x, y, z) + \frac{1-\nu}{1+r} \sum_{x' \in X} \sum_{z' \in Z} [(1 - (1-\delta)p_{P>B}^o) B^s(x', z', \Gamma') + (1-\delta)p_{P>B}^o \\
&\times \{ \sum_{\tilde{y}' \in Y} s_1 f(z', \Gamma') \{ P(x', y, z', \Gamma') + \beta \max [P(x', \tilde{y}', z', \Gamma') - P(x', y, z', \Gamma'), 0] \} g(\tilde{y}') \} \\
&+ (1 - s_1 f(z', \Gamma')) P(x', y, z', \Gamma') \} \pi_{xe}(x, x') \pi(z, z') \quad (10)
\end{aligned}$$

where \tilde{y}' denotes the match quality of the poaching firm and where the indicator for non-separation is:

$$p_{P \geq B}^o = \mathbf{1} \{ P^s(x', y, z', \Gamma') \geq B^s(x', z', \Gamma') \}.$$

Here, B^s is the value when unemployed and P^s is the value of the match at the separation stage, respectively. The first term in (10) is the flow output, the second term the value when the match separates tomorrow, the third term the value when receiving a poaching offer tomorrow and the last term the value when not receiving a poaching offer tomorrow. Also note that, regardless of what happens tomorrow, human capital while employed today increases from x to x' with probability $\pi_{xe}(x, x')$.

We also need to compute the values at the separation stage. The value for an unemployed worker at the separation stage is:

$$\begin{aligned}
B^s(x, z, \Gamma) &= (1 - f(z, \Gamma)) B(x, z, \Gamma) \\
&+ \sum_{\tilde{y} \in Y} f(z, \Gamma) [B(x, z, \Gamma) + \beta \max \{ P(x, \tilde{y}, z, \Gamma) - B(x, z, \Gamma), 0 \}] g(\tilde{y}). \quad (11)
\end{aligned}$$

The corresponding match value at the separation stage is:

$$\begin{aligned}
P^s(x, y, z, \Gamma) &= (1 - s_1 f(z, \Gamma)) P(x, y, z, \Gamma) \\
&+ \sum_{\tilde{y} \in Y} s_1 f(z, \Gamma) [P(x, y, z, \Gamma) + \beta \max \{ P(x, \tilde{y}, z, \Gamma) - P(x, y, z, \Gamma), 0 \}] g(\tilde{y}). \quad (12)
\end{aligned}$$

Then, we can simply define the surplus of a match at the matching stage as:

$$S(x, y, z, \Gamma) = P(x, y, z, \Gamma) - B(x, z, \Gamma) \quad (13)$$

and the surplus of a match at the separation stage as:

$$S^s(x, y, z, \Gamma) = P^s(x, y, z, \Gamma) - B^s(x, z, \Gamma). \quad (14)$$

Recalling that workers receive a value corresponding to their outside option plus a share β of the surplus of the match, the expected value of a new match for a firm is:

$$\begin{aligned}
J(z, \Gamma) &= \frac{1}{L} \sum_{x \in X} \sum_{y \in Y} u^s(x, z) \max\{(1 - \beta) S(x, y, z, \Gamma), 0\} g(y) \\
&\quad + \frac{s_1}{L} \sum_{x \in X} \sum_{y \in Y} \sum_{\tilde{y} \in Y} h^s(x, \tilde{y}, z) \max\{(1 - \beta) (S(x, y, z, \Gamma) - S(x, \tilde{y}, z, \Gamma)), 0\} g(y).
\end{aligned} \tag{15}$$

Note that the match-specific productivity, y , is observed when the firm meets a worker after the vacancy has been posted.⁶ The first term in (15) refers to expected surplus from recruiting out of the pool of unemployed (u^s), and the second term refers to expected surplus from recruiting from employed workers (h^s). As can be seen from (15), the distribution of unemployed workers across human capital and the distribution of matches over human capital and match-specific productivity is the endogenous aggregate state. Hence, Γ can be written as a function of L and the two terms within the summations in (15); see Appendix A.2.1 for details on how we implement this.

Let us here mention a computational aspect of the model. Solving the model is non-trivial because current values (9) and (10) depend on the probability of receiving a job offer the next period. This, in turn, depends on the next period's labor market tightness. A key determinant for the next period's tightness is the expected value of a new match to a firm in the next period, i.e., $J(z', \Gamma')$. This depends on the next period distribution of workers and firms. Fortunately, as argued in the previous paragraph, three moments fully capture the implications of this large-dimensional object. We then use a Krusell and Smith (1998)-like algorithm to let these three moments summarize and predict the labor market tightness, thereby enabling us to solve the model. For details on the solution algorithm, see Appendix A.2.

2.5 Distributional dynamics

For a new match to be formed, two conditions are required: the two parties must meet according to the meeting function (5) and the match must be an improvement over the status quo (the current match or unemployment). The unemployment distribution $u(x, z)$ after matching accordingly is:

$$u(x, z) = u^s(x, z) \left(1 - \frac{M}{L} \sum_{y \in Y} \mathbf{1}\{S(x, y, z, \Gamma) \geq 0\} g(y) \right). \tag{16}$$

⁶This assumption substantially simplifies the computation of the equilibrium.

The corresponding expression for the distribution of matches, i.e., $h(x, y, z)$, is:

$$\begin{aligned}
h(x, y, z) &= h^s(x, y, z) + \underbrace{u^s(x, z) \frac{M}{L} \mathbf{1}\{S(x, y, z, \Gamma) \geq 0\}}_{\text{mass hired from unemployment}} g(y) \\
&\quad - \underbrace{h^s(x, y, z) s_1 \frac{M}{L} \sum_{\tilde{y} \in Y} \mathbf{1}\{S(x, \tilde{y}, z, \Gamma) > S(x, y, z, \Gamma)\}}_{\text{mass lost to more productive matches}} g(\tilde{y}) \\
&\quad + \underbrace{s_1 \frac{M}{L} \sum_{\tilde{y} \in Y} h^s(x, \tilde{y}, z) \mathbf{1}\{S(x, y, z, \Gamma) > S(x, \tilde{y}, z, \Gamma)\}}_{\text{mass poached from less productive matches}} g(y). \tag{17}
\end{aligned}$$

2.6 Wage determination and worker values

Let $W(w, x, y, z, \Gamma)$ denote the present value to a worker with human capital x in a match with productivity y , wage w and aggregate productivity z . These worker values are determined according to the bargaining protocol in CPVR and are detailed as follows. Denote the renegotiated wage by w' . Workers hired out of unemployment receive the wage w' such that their value is equal to the value of unemployment plus a share β of the match surplus:

$$W(w', x, y, z, \Gamma) = B(x, z, \Gamma) + \beta S(x, y, z, \Gamma). \tag{18}$$

For employed workers who have received a poaching offer, the bargaining protocol implies that these workers receive a present value $W(w', x, y, z, \Gamma)$ equal to the value of the second-best match that they have encountered during a spell of continuous employment plus a share β of the difference in surplus between the best and second-best match. Formally, if a worker of type x employed at a firm of type y meets a firm of type \tilde{y} then, if $S(x, y, z, \Gamma) < S(x, \tilde{y}, z, \Gamma)$, the worker switches to the new firm and gets the wage w' satisfying

$$W(w', x, \tilde{y}, z, \Gamma) = P(x, y, z, \Gamma) + \beta [S(x, \tilde{y}, z, \Gamma) - S(x, y, z, \Gamma)]. \tag{19}$$

If, instead, $S(x, y, z, \Gamma) \geq S(x, \tilde{y}, z, \Gamma)$, the worker remains in his current match and gets a wage w' that satisfies:

$$W(w', x, y, z, \Gamma) = \max \{P(x, \tilde{y}, z, \Gamma) + \beta [S(x, y, z, \Gamma) - S(x, \tilde{y}, z, \Gamma)], W(w, x, y, z, \Gamma)\}. \tag{20}$$

Note that, in case the value at the current wage is higher than the one implied by the outside option, the wage is unchanged.

Wages for workers who do not receive poaching offers can also be rebargained, as aggregate or

idiosyncratic shocks might affect the various values. First, if the wage is such that it implies a worker value that is larger than the match value, then the match would break down unless there is renegotiation. Hence, the wage is then set so that $W(w', x, y, z, \Gamma) = P(x, y, z, \Gamma)$. Second, if the wage is such that the worker value is lower than $B(x, z, \Gamma) + \beta S(x, y, z, \Gamma)$, the worker can ask for a renegotiation with unemployment as the outside option. Hence, the wage is then set so that $W(w', x, y, z, \Gamma) = B(x, z, \Gamma) + \beta S(x, y, z, \Gamma)$. Finally, the current wage w is unchanged when:

$$B(x, z, \Gamma) + \beta S(x, y, z, \Gamma) \leq W(w, x, y, z, \Gamma) \leq P(x, y, z, \Gamma). \quad (21)$$

To solve for wages, we compute the value for a worker earning w today, given that future values are (partially) determined by (18)-(21). An employed worker earning the wage w in the current period faces four possibilities in the next period: i) staying employed and not meeting any new firm, ii) staying employed and receiving a successful poaching offer and switching jobs, iii) staying employed and receiving an unsuccessful poaching offer (and staying in the old job) and iv) separating. Note that, if the worker becomes separated in the next period he still has a chance to find a new job within the period. Imposing an interior solution for M , $M = \alpha L^\omega V^{1-\omega}$ and using the definition of q , the probability of meeting a new firm for an employed worker is $s_1 f(z', \Gamma')$. Then, given the wage, w , the worker value (at the matching stage) is:

$$\begin{aligned} W(w, x, y, z, \Gamma) = & w + \frac{1-\nu}{1+r} \sum_{x' \in X} \sum_{z' \in Z} [(1-s') \{(1-s_1 f(z', \Gamma')) W'_{np} \\ & + s_1 f(z', \Gamma') \sum_{\tilde{y} \in Y} (p_{\tilde{y}>y}^o W'_{p, \tilde{y}>y} + (1-p_{\tilde{y}>y}^o) W'_{p, \tilde{y} \leq y}) g(\tilde{y})\} \\ & + s' \left(B(x', z', \Gamma') + f(z', \Gamma') \sum_{y' \in Y} \beta S(x', y', z', \Gamma') g(y') \right)] \pi_{xe}(x, x') \pi(z, z'), \end{aligned} \quad (22)$$

where

$$\begin{aligned} s' &= (\mathbf{1} \{S(x', y, z') < 0\} + \delta \mathbf{1} \{S(x', y, z', \Gamma') \geq 0\}) \\ W'_{np} &= \min \{P(x', y, z', \Gamma'), \max \{W(w, x', y, z', \Gamma'), B(x', z', \Gamma') + \beta S(x', y, z', \Gamma')\}\} \\ p_{\tilde{y}>y}^o &= \mathbf{1} \{S(x', \tilde{y}, z', \Gamma') > S(x', y, z', \Gamma')\} \\ W'_{p, \tilde{y}>y} &= P(x', y, z', \Gamma') + \beta [S(x', \tilde{y}, z', \Gamma') - S(x', y, z', \Gamma')] \\ W'_{p, \tilde{y} \leq y} &= \max \{P(x', \tilde{y}, z', \Gamma') + \beta [S(x', y, z', \Gamma') - S(x', \tilde{y}, z', \Gamma')], W(w, x', y, z', \Gamma')\}, \end{aligned}$$

where s' denotes separations, W'_{np} the value when not receiving a poaching offer, $p_{\tilde{y}>y}^o$ a successful poaching offer, $W'_{p, \tilde{y}>y}$ the value of a successful poaching offer and $W'_{p, \tilde{y} \leq y}$ the value of an unsuccessful

poaching offer.

2.7 Wage distribution

When determining the wage distribution, it follows from the description of the wage setting above that the current wage of the worker is a state variable. The distribution of matches over w , x and y after separations is:

$$h^{s,w}(w, x, y, z) = (1 - \delta)(1 - \nu) \sum_{x_{-1} \in X} \mathbf{1}\{S^s(x, y, z, \Gamma) \geq 0\} h^w(w, x_{-1}, y, z_{-1}) \pi_{xe}(x_{-1}, x). \quad (23)$$

Analogously to (17) in section 2.5, we define $h^w(w, x, y, z)$, i.e., the distribution after matching and wage rebargaining; see Appendix A.1.

3 Calibration

3.1 Distributions and shock processes

The log of the exogenous part of TFP, z , follows an AR(1) process approximated by a Markov chain. The log of match productivity, $g(y)$, is normally distributed and its mean value is normalized to 0.5. The number of gridpoints for x , y and z are 10, 8 and 5, respectively.⁷ The wage grid contains 15 points and is chosen separately for each parameter vector so as to only cover the relevant wage interval.⁸ In constructing the grid for human capital, x , we, as e.g., Jarosch (2015), follow Ljungqvist and Sargent (1998, 2008) in using an equal-spaced grid and in setting the ratio between the maximum and minimum value of x to 2. The structure of the transition matrices $\pi_{xe}(x, x')$ and $\pi_{xu}(x, x')$ for human capital also closely follows Ljungqvist and Sargent. Abstracting from the bounds, the probability of an employed worker to increase his human capital by one gridpoint is x_{up} and the probability for an unemployed worker to decrease his human capital by one gridpoint is x_{dn} . With the reciprocal probabilities, the human capital of a worker is unchanged.

3.2 Calibration approach

The frequency of the model is monthly. We calibrate the model based on U.S. data. Parameters whose values are well established in the literature or can be set based on model-independent empirical evidence are set outside the model. Table 1 documents these parameter values and their sources.

⁷For z , we use Tauchen and Hussey's (1991) discretization of AR(1) processes with optimal weights from Flodén (2008). This algorithm has been shown by Flodén (2008) to also be accurate for processes with high persistence.

⁸The coarseness of the wage grid is less restrictive than it seems, as we map each "off-the-grid" wage to the two nearest grid points using the inverse of the distance to the grid point as weight. Furthermore, the wage grid has no impact on the allocations in the model.

Table 1: Parameters set outside the model

	Explanation	Value	Source
ω	Matching function elasticity	0.5	Pissarides (2009)
δ	Exogenous match separation rate	0.030	Fujita-Ramey (2009)
c_0	Vacancy posting cost	0.06375	Fujita-Ramey (2012)
ν	Retirement rate	$1/(40 * 12)$	40-year work-life
ρ	TFP shock persistence	0.960	Hagedorn-Manovskii
r	Interest rate	$1.05^{1/12} - 1$	Annual r of 5%

The meeting function elasticity, ω , is set in line with the convention in the literature. The exogenous match separation rate, δ , is equal to the mean E2U transition rate reported by Fujita and Ramey (2009), adjusted for workers finding a new job the same month as they lost the old job.⁹ We set the vacancy posting cost c_0 along the lines for Fujita and Ramey (2012) who refer to evidence that vacancy costs are 6.7 hours per week posted. We set the retirement (or death) rate to match an average work-life of 40 years, as e.g. Huckfeldt (2016). To compute the persistence of the AR process for TFP, we follow along the lines of Hagedorn and Manovskii (2008). Specifically, we simulate a monthly Markov chain to match a quarterly autocorrelation of (HP-filtered) log labor productivity of 0.765. Finally, we set r to yield an annualized interest rate of 5% as in LR.

Table 2: Parameters obtained by moment-matching

Parameter	Explanation	Value	Main identifying moment
α	Matching function productivity	0.686	U2E transition rate, mean
s_1	Relative search intensity of employed	0.426	J2J transition rate, mean
x_{up}	Human capital gain, probability	0.0427	Return to experience
b_0	Unemployment payoff	0.374	Unemployment, std.dev.
β	Bargaining strength of workers	0.848	Wage elasticity wrt productivity
σ_y	Match-specific productivity dispersion	0.259	Wage disp: Mean-min ratio
$100\sigma_z$	TFP shock std.dev.	0.698	GDP, std.dev.

The remaining parameters of our model are calibrated jointly to match key moments. For simplicity, and in line with most of the literature, flow payoff from unemployment is $b(x, z) = b_0$, i.e. invariant of aggregate productivity and human capital. Table 2 documents the 7 calibrated parameters and the 7 moments matched, including the main identifying moment for each parameter. We minimize the squared percentage deviation between model and data moments. Let us now motivate the choice of moments. Note first, that since we are interested in the cost of business cycles from a mechanism driven by unemployment volatility, it is important to match GDP and unemployment volatility. Turning to identification, the model parameters are jointly estimated, but some moments

⁹The latter implies that the separation rate exceeds the E2U rate by a factor of $1/(1-\text{job finding rate})$. By using Fujita and Ramey's number for E2U transitions, we control for the fact that empirically, but not in our model, workers flow in and out of the labor force.

are more informative about certain parameters. The mean transition rate from unemployment to employment is informative about the matching function productivity α . The job-to-job transition rate is informative about the relative search intensity of employed workers s_1 . Return to experience, measured as the average percentage wage increase while employed, is informative about on-the-job accumulation of human capital, x_{up} .¹⁰ Unemployment volatility is informative about the unemployment payoff parameter, b_0 . As pointed out by Hagedorn and Manovskii (2008), wage elasticity with respect to labor productivity is informative regarding worker bargaining strength, β . Wage dispersion is informative about the dispersion of match-specific productivity, σ_y . Finally, the volatility of GDP and unemployment are both informative about the standard deviation of the aggregate productivity process.

Table 3: Data moments and matched model moments

Moment	Data source	Target value (data)	Model value
U2E transition rate, mean	Fujita-Ramey (2009)	0.340	0.357
J2J transition rate, mean	Moscarini-Thompson	0.0320	0.0290
Unemployment, std.dev.	BLS 1980-2010	0.107	0.0973
GDP, std.dev.	BEA 1980-2010	0.0136	0.0136
Wage disp: Mean-min ratio	Hornstein et al.	1.50	1.70
Wage elasticity wrt productivity	Hagedorn-Manovskii	0.449	0.445
Return to experience	Buchinsky et al.	0.0548	0.0518

Notes: U2E and J2J transition rates are at a monthly frequency. Unemployment is a quarterly mean of a monthly series. This variable, as well as GDP, labor productivity and aggregate wages, have been logged and HP-filtered with $\lambda = 1,600$, both in the data and the model.

Let us comment on the more unusual data used. The relevant measure of wage dispersion for our model is “residual” wage dispersion, i.e. controlling for heterogeneity not present in the model, such as education, sex, race etc. We take the mean-min ratio (capturing the minimum by the 10th wage percentile) from Hornstein, Krusell and Violante (2007) as our measure of wage dispersion. We use their preferred measure of 1.50, which is an average of their ratios from census, OES and PSID data. Similarly to Kehoe et al. (2015) we use estimates from Buchinsky et al. (2010) to obtain the “return to experience”. Specifically, from Buchinsky’s estimated coefficients we obtain the marginal return to experience of a worker in his third year of employment. We then match that to the wage increase

¹⁰As in Jarosch (2015), we impose a relationship between x_{up} and x_{dn} such that the number of increases in human capital roughly equals the number of decreases to minimize bunching at end-points of the human capital grid X . In particular, letting u^{tot} denote the (implicitly, through the mean values of E2U and U2E) targeted value of unemployment, we impose $(1 - \nu) x_{up} (1 - u^{tot}) \Delta x = (1 - \nu) x_{dn} u^{tot} \Delta x + \nu (\bar{x} - \underline{x})$ where Δx is the distance between two gridpoints and \bar{x} represents average human capital for dying workers. For computational reasons, we set \bar{x} to the midpoint of the grid. Furthermore \underline{x} is the lower bound of the grid, representing the human capital of newly born workers. This implies $x_{dn} = \left(x_{up} - \frac{\nu}{1-\nu} \frac{[\bar{x}-\underline{x}]}{(1-u^{tot})\Delta x} \right) \frac{1-u^{tot}}{u^{tot}}$. There will still be some upward drift, and thereby upper end-point bunching, in the human capital distribution if an above-proportional fraction of the unemployed are at the lower bound of the human capital grid, unless this is offset by the analogous force of above-proportional fraction of employed workers at the upper bound.

of workers in the model who works for three years for the same employer. We can thereby keep the match-specific productivity fixed and obtain a clean measure of the effect of human capital on wages.

4 Results

4.1 Targeted moments and the parameter estimates

The moment-matching exercise can be evaluated by comparing the last two columns in Table 3. The model is able to fit most of these moments well, with less than 10 percent deviation for all but one moment, wage dispersion.

It might appear surprising that we need to calibrate the volatility of (the exogenous part of) TFP, but this is necessary since the model has internal amplification and propagation of the exogenous TFP shocks, as the distribution of human capital of workers, the productivity of matches and sorting between workers and jobs varies over the cycle. All of this implies that measured TFP in our model is a combination of exogenous TFP and endogenous propagation.¹¹

The above moment-matching exercise determines the 7 parameters in Table 2. The value for s_1 in Table 2 indicates that employed workers meet prospective employers slightly below half as often as unemployed workers. We follow LR and report the replacement ratio for unemployed workers as a fraction of the output of the best possible match. The value of b_0 implies that this ratio is 0.600, averaged over the human capital values. We find that that worker bargaining strength is fairly high, 0.848.

Given their centrality for our mechanism, we report and comment in more detail on our estimates of the parameters determining the human capital dynamics. The estimated Markov transition probability ($x_{up} = 0.0427$) imply that the expected monthly human capital increase for an employed worker is 0.207 percent, while the expected decrease when unemployed is 1.41 percent (for $x_{dn} = 0.557$).¹²

We know of only one direct measure in the literature of general human capital loss while non-employed: Edin and Gustavsson (2008). They use a Swedish panel of individual level data that includes test results on labor market-relevant general skills and information about employment status between test dates. First, they find that time-out-of-work (compared to employment) implies skill

¹¹One could potentially also calibrate the persistence of exogenous TFP jointly with the 7 parameters in Table 2 to match e.g., the persistence of GDP. However, to reduce computational complexity we calibrate this parameter as outlined above. Moreover, the persistence of GDP turns out to be fairly well matched in our calibration; it is 0.801 in the model compared to 0.867 in the data.

¹²This value takes into account the distribution of employed and unemployed workers across the human capital grid, including the effects of the bounds of the human capital grid.

loss, significant at the 1% level. Second, this skill loss appears to be linear in time out-of-work. Third, the speed of skill loss is substantial; being out-of-work for a year implies losing skills equivalent to 0.7 years of schooling.

The human capital dynamics can be compared to estimates in models broadly similar to ours.¹³ Huckfeldt (2016) reports a 0.330 percent expected monthly human capital increase for workers in skill-intensive jobs (0.220 percent in skill-neutral jobs). For unemployed workers Huckfeldt obtains a gradual human capital decrease of 1.13 percent per month.¹⁴ Jarosch (2015) reports only the monthly human capital Markov transitions probabilities: 0.0141 for employed and 0.131 for unemployed. In Jarosch (2015), for an employed worker with the mid-point of human capital, this implies an expected increase of 0.134 percent, and for the unemployed worker with the mid-point of human capital, it implies a 1.25 percent decrease. To sum up this comparison to the literature, our human capital accumulation for employed workers is in between the estimates of Huckfeldt (2016) and Jarosch (2015), while for unemployed workers our value is about as large as their estimates.

4.2 Welfare measure

As is standard in the cost of business cycle literature since Lucas (1987), we report the amount of consumption agents are willing to forego to eliminate business cycles. Specifically to our model, the linearity of utility in consumption makes welfare calculations straightforward, since then the flow of aggregate welfare is proportional to aggregate consumption. To compute market consumption, we deduct vacancy posting costs from GDP. Note that one may interpret the unemployment payoff, b , in two ways, which has different welfare implications. In the first interpretation, b is home production (or equivalently, from a welfare perspective, utility of leisure) in which case the welfare relevant quantity is the sum of market consumption and the unemployment payoff. In the second interpretation, b is a pecuniary transfer with no direct effect on aggregate utility. We report results for both interpretations.¹⁵

¹³First, there is an older empirical literature that attributes all wage loss when re-employed after an unemployment spell to human capital loss and furthermore assumes that the wage equals marginal product of labor. This is not consistent with our model so we can not use that literature for calibration or straight comparison. Second, some papers look at the effect on wages of an additional month of unemployment. The estimates in Neal (1995) imply that an additional month of unemployment reduces the re-employment wage by 1.5%, which, under the assumption that the wage equals marginal product of labor, is very much in line with the results here. Recent results by Schmieder et al. (2016) shows that re-employment wages decrease by 0.8% per (additional) month unemployed. This is somewhat lower than our result, but reasonably well in line if we think that there is some surplus sharing so that wages decrease less than human capital for an additional month of unemployment.

¹⁴The comparison of skill losses during unemployment to Huckfeldt's results is clouded by the fact that, in contrast to our model, he allows for both gradual and sudden loss of human capital during unemployment. Our (gradual) human capital loss estimates for unemployed workers will therefore tend to be higher than his.

¹⁵There is also an intermediate case where b consists of both home production and transfers. The welfare cost of eliminating aggregate volatility generated by our mechanism will then fall between these two cases.

4.3 Results for cost of business cycles

Our main exercise is to compute the consequences for welfare, GDP and employment of eliminating aggregate volatility. As documented in Table 4, we find that in our model the elimination of aggregate volatility increases steady state GDP by a substantial amount, 1.45 percent. This also has consequences for steady state consumption and welfare, which increase by 0.52-1.49 percent depending on the interpretation of the unemployment payoff. As we will document below, these fairly large effects are due to the positive relationship between employment and human capital accumulation.

From an accounting perspective, the increase in GDP can be decomposed into the increase in employment and the change in the average level of human capital of employed workers, $E(x \times h(\cdot))$, respectively.¹⁶ Of these two, the increase in employment accounts for the vast majority. To understand the effects of human capital on employment, recall from (15) that job creation is affected by the human capital of both employed and unemployed workers. In our calibration, the effects through the unemployed dominates. This is partly due to that the average levels of human capital for the unemployed changes more; $E(x \times u(\cdot))$ increases by 4.36 percent while $E(x \times h(\cdot))$ increases by 0.18 percent. In addition, job creation is much more sensitive to changes in human capital of the unemployed. Specifically, the elasticity of $J(z, \Gamma)$ with respect to $E(x \times u(\cdot))$ is 1.27 while the elasticity of $J(z, \Gamma)$ with respect to $E(x \times h(\cdot))$ is 0.39. It may be surprising that the change in $E(x \times h(\cdot))$ is so moderate. However, the reason is that the composition of the employed workers is affected by the elimination of business cycles. Specifically, in the absence of aggregate volatility, the positive effect that higher employment has on human capital is counteracted by the tendency that firms tend to hire a larger fraction of workers with low human capital.

Table 4: Steady state effects of eliminating business cycles (in percent)

	Baseline	No human capital dynamics
Welfare, b transfer, (GDP-vacancy cost)	1.49	0.26
Welfare, b home prod, (GDP-vacancy costs+ $b * u$)	0.52	0.02
GDP	1.45	0.25
Employment	1.34	0.34
$E(x \times h(\cdot))$	0.18	0
$E(x \times u(\cdot))$	4.36	0

¹⁶Although negligible for our exercise, there are other factors affecting average productivity than human capital. Examples include the change in the average level of match-specific productivity, $E(y \times h(\cdot))$, and the changed degree of sorting between workers and firms (as well as the covariation between any of these objects with the cycle).

4.3.1 The importance of human capital dynamics

Let us now quantify the importance of the change in the human capital distribution for the cost of business cycles. To do this we perform a counterfactual exercise where we keep the human capital distribution of the population (i.e. combining employed and unemployed workers) fixed when we remove the aggregate volatility. All other aspects of the computation is the same as in the baseline exercise.¹⁷ The last column of Table 4 confirms the importance of learning on-the-job, as the version of our model without human capital dynamics implies that aggregate fluctuations have negligible effects on the average level of welfare, GDP and employment.

4.3.2 Accounting for the transition

We now compute the welfare consequences of eliminating aggregate volatility taking the transition dynamics into account. This is unique in the macro-labor literature on the cost of business cycles; previous studies have only compared steady state quantities. As reported in Table 5, we find that in our model, the elimination of aggregate volatility when taking the transition into account, increases welfare by 0.20-1.09 percent depending on the interpretation of the unemployment payoff.¹⁸ We note that the welfare gains from removing business cycles are lower when accounting for the transition than when simply comparing steady states. The gains when accounting for the transition are lower for two reasons: discounting of the increased future consumption and the extra vacancy posting costs related to the increase in employment along the transition path. Note also that the transition to the non-stochastic steady state is reasonably fast. The half-time of the transition of GDP is 4.5 years.

Table 5: Welfare effects of eliminating business cycles (in percent)

Welfare, b transfer	1.09
Welfare, b home prod	0.20

4.3.3 Robustness

Two key determinants of the cost of business cycles in our model are i) how sensitive the human capital distribution is to the change in (un)employment, and ii) how sensitive job creation is to changes in the human capital distribution of both unemployed and employed workers. An important factor affecting

¹⁷We fix the human capital distribution by setting $x_{up} = \nu = 0$ and assume that it is given by the average distribution in the baseline calibration with aggregate volatility. We also keep the incentives for job creation and destruction unchanged, i.e. S and B are computed with the baseline human capital parameters.

¹⁸We compute welfare when taking the transition into account in the following way. First, we simulate the economy with aggregate volatility for several thousand periods. We then draw 1000 starting points for the transition from this simulation and compute welfare in each of these starting points, given that productivity is constant at its mean value for all future periods. Finally, we calculate the mean across the 1000 transitions.

the sensitivity of the human capital distribution is the range of values that human capital can take and an important factor affecting the sensitivity of job creation to human capital is the bargaining strength of workers.

Thus, to judge the robustness of the results we re-calibrate it under alternative assumptions on the human capital distribution and the bargaining power and report the steady state welfare, GDP and employment cost of business cycles in Table 6. First, we document what the cost of business cycles is when allowing for a wider range of values for human capital. Recall that in our main calibration we have followed Ljungqvist and Sargent (1998, 2008) and assumed that the ratio between the highest and the lowest human capital value is 2. Huckfeldt (2016) instead finds a ratio of 15.25. Here we illustrate the effects of changing the assumption regarding the human capital range in the direction of Huckfeldt by assuming that the maximum ratio of human capital is 4. We then re-calibrate the model by matching the same moments as above in Table 3. We find that eliminating aggregate volatility lead to an increase of welfare and GDP of 0.94-1.94 and 1.89 percent, respectively. In other words, the cost of business cycles increase substantially. The main difference compared to our baseline calibration is that GDP increases much more than employment indicating that the wider human capital range generated a larger increase in average productivity from the elimination of business cycles. The result of this exercise implies that the cost of business cycles might be substantially higher than what we obtain when using the quite conservative parametrization of the human capital range from Ljungqvist and Sargent (1998, 2008).

Second, we explore the sensitivity of our results to the bargaining strength of workers. In particular, we fix the bargaining power at 0.50, as is commonly done in the literature that, differently from our setup, considers Nash bargaining with unemployment as the (only) outside option of the worker. We then re-calibrate the model by matching the same moments as above in Table 3, except the elasticity of wages, that was used to identify bargaining power in the baseline calibration. We find that when $\beta = 0.50$, the elimination of business cycles have somewhat larger effects on all variables compared to our baseline calibration.

Table 6: Steady state effects of eliminating business cycles under alternative assumptions (in percent)

Model version	Welfare, b transfer	Welfare, b home prod	GDP	Employment
Baseline	1.49	0.52	1.45	1.34
Wider range of human capital	1.94	0.94	1.89	1.43
$\beta = 0.50$	1.78	0.56	1.83	1.42

5 Conclusions

A central question in macroeconomics is how large the welfare costs of business cycles are. We show that cyclical variation in unemployment reduces aggregate welfare in a labor market search model with general human capital dynamics since it drives down the level of employment, output and consumption. The key mechanism of the paper concerns learning on-the-job and skill loss during unemployment and is as follows. Empirically, the Beveridge correlation is negative, i.e., vacancies and unemployment are negatively correlated. This, in turn, means that business cycles tend to reduce the average number of matches and hence employment through the matching function. Then, since learning on-the-job and skill loss during unemployment implies that human capital is increasing in the employment rate, it follows that aggregate volatility reduces human capital. This, in turn, reduces incentives to post vacancies, further reducing employment. We find that the steady state output and welfare gains from eliminating business cycles are large - they amount to 1.45 percent and 0.52-1.49 percent, respectively. The alternative parameter assumptions explored indicate that the cost of business cycles might be higher than this. We also show that human capital dynamics is pivotal for the results - if we disable this mechanism in our model, the implied gains in employment, GDP and welfare from eliminating business cycles are negligible.

To conclude, let us briefly discuss some broader implications of our results. In our model, there is only one type of aggregate shock. If we view this shock as a “catch-all” for any variation in firm revenues including effects of fiscal and monetary policy, we can draw interesting policy conclusions. In particular, a policy that successfully stabilizes unemployment (or job-finding rates) raises the average level of output. For this reason, our paper rationalizes an unemployment stabilization mandate for monetary and fiscal policy. In this sense we reach the same conclusion as Berger et al. (2016) and Galí (2016) but for a very different reason. Berger et al.’s argument is about unemployment stabilization reducing idiosyncratic risk related to layoffs, while Galí’s mechanism is about hysteresis due to insider-outsider dynamics. Our mechanism is about unemployment stabilization leading to a higher average level of output, thereby more closely related to the argument by Summers (2015) that stabilization policy can have major effects on average levels of output over periods of decades.

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A Appendix

A.1 Employment transitions

When accounting for the wage distribution, the employment transition follows:

$$\begin{aligned}
h^w(w^*, x, y, z) = & \\
& \underbrace{h^{s,w}(w^*, x, y, z) - h^{s,w}(w^*, x, y, z) s_1 \frac{M}{L} \sum_{\tilde{y} \in Y} p_{\tilde{y} > y}^o g(\tilde{y})}_{\text{mass lost to more productive matches}} \\
& - \underbrace{h^{s,w}(w^*, x, y, z) s_1 \frac{M}{L} \sum_{\tilde{y} \in Y} \mathbf{1}\{P\beta(x, \tilde{y}, y, z, \Gamma) > W(w^*, x, y, z, \Gamma)\} (1 - p_{\tilde{y} > y}^o) g(\tilde{y})}_{\text{mass lost to higher wage offers from less productive matches}} \\
& + \underbrace{s_1 \frac{M}{L} \sum_{\tilde{y} \in Y} \sum_{\tilde{w} \in W^{grid}} h^{s,w}(\tilde{w}, x, y, z) \mathbf{1}\{w(\tilde{w}, x, y, z, \Gamma) = w^*\} (1 - p_{\tilde{y} > y}^o) g(\tilde{y})}_{\text{mass gained from increased wage due to offers from less productive matches}} \\
& + \underbrace{s_1 \frac{M}{L} g(y) \sum_{\tilde{y} \in Y} h^s(x, \tilde{y}) \mathbf{1}\{W(w^*, x, y, z, \Gamma) = P(x, \tilde{y}, z, \Gamma) + \beta[S(x, y, z, \Gamma) - S(x, \tilde{y}, z, \Gamma)]\} p_{y > \tilde{y}}^o}_{\text{mass poached from less productive matches}} \\
& - \underbrace{h^{s,w}(w^*, x, y, z) \mathbf{1}\{W(w^*, x, y, z, \Gamma) \notin BS(x, y, z, \Gamma)\}}_{\text{mass lost due to being outside bargaining set}} \tag{24} \\
& + \underbrace{\sum_{\tilde{w} \in W^{grid}} h^{s,w}(\tilde{w}, x, y, z) \mathbf{1}\{w(\tilde{w}, x, y, z, \Gamma) = w^*\} \mathbf{1}\{W(\tilde{w}, x, y, z, \Gamma) \notin BS(x, y, z, \Gamma)\}}_{\text{mass gained from other wages being outside bargaining set}} \\
& + \underbrace{\frac{M}{L} u^s(x) g(y) S_{xyz} \mathbf{1}\{W(w^*, x, y, z, \Gamma) = B(x, z, \Gamma) + \beta S(x, y, z, \Gamma)\}}_{\text{mass hired from unemployment}}
\end{aligned}$$

where W^{grid} is the wage grid and

$$\begin{aligned}
p_{\tilde{y} > y}^o &\equiv \mathbf{1}\{P(x, \tilde{y}, z, \Gamma) > P(x, y, z, \Gamma)\} \\
P\beta(x, \tilde{y}, y, z, \Gamma) &= P(x, \tilde{y}, z, \Gamma) + \beta[S(x, y, z, \Gamma) - S(x, \tilde{y}, z, \Gamma)] \\
p_{y > \tilde{y}}^o &\equiv \mathbf{1}\{P(x, y, z, \Gamma) > P(x, \tilde{y}, z, \Gamma)\} \\
BS(x, y, z, \Gamma) &= [B(x, z, \Gamma) + \beta S(x, y, z, \Gamma), P(x, y, z, \Gamma)] \\
S_{xyz} &\equiv \mathbf{1}\{S(x, y, z, \Gamma) \geq 0\}
\end{aligned}$$

A.2 Solution algorithm

A.2.1 Preliminaries

As can be seen from (9) and (10), the values B and P depend on Γ' through the job finding rate, and thereby the entire expected next period distribution of matches across x and y and unemployed workers distribution over x . The challenge is to reduce the dimensionality of the distributions Γ' to something manageable. The key to our algorithm is to note that all influence of the endogenous distributions goes through the next period labor market tightness, θ' . In addition, according to (7) labor market tightness is a function only of L in (3) and J in (15). Hence, we can write θ as a function of the three moments that make up (3) and (15); $\theta = \Theta(m_1, m_2, m_3; z)$. In particular, noting that $\sum_{x \in X} \sum_{y \in Y} h^s(x, y, z) = 1 - \sum_{x \in X} u^s(x, z)$ and accordingly $L_t \equiv \sum_{x \in X} u^s(x, z) + s_1(1 - \sum_{x \in X} u^s(x, z))$ we set

$$m_1 = \sum_{x \in X} u^s(x, z). \quad (25)$$

Given that L_t can be computed using m_1 , equation (15) implies that J is fully determined by m_1 and the following additional two terms:

$$m_2 = \sum_{x \in X} \sum_{y \in Y} u^s(x, z) \max\{S(x, y, z, \Gamma), 0\} g(y) \quad (26)$$

and

$$m_3 = \sum_{x \in X} \sum_{y \in Y} \sum_{\tilde{y} \in Y} h^s(x, \tilde{y}, z) \max\{S(x, y, z, \Gamma) - S(x, \tilde{y}, z, \Gamma), 0\} g(y). \quad (27)$$

To compute next period values of these moments we assume a linear relationship to today's moments. Thus, we write

$$m'_m = H_m(m_1, m_2, m_3, z'). \quad (28)$$

Note that, similarly to LR, we can compute the evolution of the distributions u^s and h^s and θ without solving for wages and worker values. However, in contrast to LR, match surpluses and the value unemployment is jointly determined with (tomorrow's) labor market tightness. Therefore we guess functions Θ and H_m for labor market tightness and the evolution of moments. We can then compute match values. Given the solution for match values we can compute the allocation for a sequence of aggregate productivity shocks and then update the guesses for Θ and H_m using standard estimation methods and iterate until convergence (see Krusell and Smith (1998)). Given the above arguments it is unsurprising that the R^2 of the function $\Theta(m_1, m_2, m_3)$ is approximately unity (≥ 0.9997). It turns out that $H_m(m_1, m_2, m_3, z')$ also has a high R^2 . In the end, we can replace the distributions in Γ' by (m_1, m_2, m_3) so that instead of (w, x, y, z, Γ) the final state vector is $(w, x, y, z; m_1, m_2, m_3)$.

We discretize m_i on a grid. We choose fewer gridpoints for m_i (2 gridpoints) than for z as m_i is quantitatively less important. With the functions Θ and H_m at hand, we solve for values B and P and then residually compute S .

A.2.2 Detailed algorithm

Step 1. Obtain the equilibrium without aggregate volatility by the following substeps:

- i) Make an initial guess for equilibrium unemployment.
 - ii) Set the parameter $x_{dn} = \left(x_{up} - \frac{\nu}{1-\nu} \frac{[\bar{x}-\underline{x}]}{(1-u^{tot})\Delta x} \right) \frac{1-u^{tot}}{u^{tot}}$ where u^{tot} is the (implicitly, through the mean values of E2U and U2E) targeted unemployment rate (i.e., $u^{tot} = 0.0556$), \bar{x} is the midpoint of the x -grid and Δx is the distance between two gridpoints. Furthermore \underline{x} is the lower bound of the grid, representing the human capital of newly born workers.
 - iii) Guess the ergodic job finding rate f .
 - iv) Use value function iteration to solve for ergodic B and P jointly. Note that the ergodic versions of B and P corresponding to expressions (9) and (10) can be written as a function of x , y , \bar{z} and f only. Then compute ergodic S along the lines of (13).
 - v) Compute the ergodic distributions for $u(x)$ and $h(x, y)$ for a fixed $z = \bar{z}$ (see below for details).
 - vi) Compute the equilibrium job finding rate f' . If f' is close to f then continue. Otherwise return to iv).
 - vii) If unemployment is too different from previous value, go back to iii).
- To obtain the ergodic distributions for $u_{t+1}(x)$ and $h_{t+1}(x, y)$ simulate above for a fixed z until convergence in these distributions.

Step 2. Draw a sequence $\{z_t\}_{t=0\dots T}$ and guess functions Θ and H_m .

Step 3. Use value function iteration to solve for $B(x, z, \Gamma)$ in (9) and $P(x, y, z, \Gamma)$ in (10) jointly, interpolating next period values over next period moments. Then compute $S(x, y, z, \Gamma)$ in (13).

Step 4. For each t , guess current moments (m_1, m_2, m_3) .

- i) Interpolate S on the moments.
- ii) Given interpolated S , we can solve for the allocation objects we are interested in:
- iii) Calculate $u_t^s(x)$ and $h_t^s(x, y)$ using (1) and (2)
- iv) Calculate L_t by aggregating over $u_t^s(x)$ and $h_t^s(x, y)$
- v) Calculate J_t using (15).
- vi) Calculate θ_t using (7)
- vii) Calculate V_t using (6)
- viii) Calculate $u_{t+1}(x)$ and $h_{t+1}(x, y)$ using (16) and employment transition (17)
- ix) Compute updated moments $(m_1^{new}, m_2^{new}, m_3^{new})$

x) If $(m_1^{new}, m_2^{new}, m_3^{new})$ is close to (m_1, m_2, m_3) we are done. Otherwise, return to i).

Step 5. Update the functions Θ' and H'_m using the regressions described in A.2.1 with the time series for m_1 , m_2 and m_3 and tightness θ . If Θ' and H'_m are close to Θ and H_m we are done. Otherwise, return to Step 3 with the new guess.

Given the sequence based on $\{z_t\}_{t=0\dots T}$ above, we use the resulting sequence of θ (after removing an initial burn-in period) to compute allocations and wages and then the sequence of h_{t+1}^w to compute relevant moments of the wage distribution along the sequence where we have followed the algorithm described in section A.2.3 to compute worker values $W(w, x, y, z, \Gamma)$ and wages $w(w, x, y, z, \Gamma)$.

A.2.3 Algorithm for determination of W and w

With the functions Θ and H_m found in section A.2.2, we solve for worker values W , noting that the state vector is $(w, x, y, z; m_1, m_2, m_3)$. The solution is obtained by value function iteration, interpolating next period values over next period moments.

Once we know the worker values W we can solve for wages w residually. This amounts to rewriting equation (22) to find the wage that yields the right value of W for the current state vector $(w, x, y, z; m_1, m_2, m_3)$ given the expected future values for the worker. In all computations related to wages we interpolate linearly over the moments.

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